Markov model of unsteady profile of normal behavior for network objects of computer systems

$$\label{eq:loss} \begin{split} & \text{Ihor Tereikovskyi}^{1[0000-0003-4621-9668]}, \text{Lyudmila Tereykovska}^{2[0000-0002-8830-0790]}\,,\\ & \text{Shynar Mussiraliyeva}^{3[0000-0001-5794-3649]}, \text{Mikola Tsiutsiura}^{2[0000-0002-1946-9242]}\,,\\ & \text{Kalamkas Bagitova}^{3[0000-0003-1587-1995]} \end{split}$$

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Abstract. The article is devoted to the ongoing scientific and applied issue on improvement of systems of detection of cyberattacks on network objects of computer systems. Detection systems are considered based on determining the tolerance of deviations of the current values of the controlled functional parameters of the computer system from the profiles of normal behavior. It is established that one of the main disadvantages of network cyberattack detection systems is the imperfection of normal behavior profiles, which are insufficiently adapted to the typical non-stationary nature of the dynamics of the controlled functional parameters. It is proposed to form non-stationary profiles of normal behavior of network objects of computer systems on the basis of multiperiodic Markov model, which allows to take into account the typical nature of the dynamics of functional parameters that reflect the state of security of network objects of computer systems. The peculiarity of the model is the modeling of each of the stationary sections of the dynamics of the functional parameter using a homogeneous Markov chain with successive transitions. It is experimentally established that the application of the developed multiperiodic model allows to increase the accuracy of forecasting the dynamics of functional parameters up to 2 times. Moreover, it is shown that the prospects for further research are associated with the development of methods for applying the solutions of the theory of spectral analysis of data to determine the significant periods of the process of changing functional parameters.

Keywords: recognition of network cyberattacks, normal behavior profile, Markov's model.

1 Introduction

The importance of improving the means of recognizing network cyberattacks is increasing due to the constant danger of cybercrime, the dynamic growth of various and diverse cyber threats, and the increase in the complexity and responsibility of computer networks (CN) [1, 5, 11]. The basis of these tools is the data analysis module, the main task of that is to generate a decision on the presence or absence of a cyberattack on the object of protection at a given time [8, 10, 16]. It is generally recognized that one of the most promising areas for improving the quality of data analysis is the use of the anomaly detection method in recognition systems. An important advantage of this method is the ability to recognize new types of attacks with unknown signatures. In this case, the work of data analyzers is based on the assumption that the sign of a network cyberattack is a deviation of the specified parameters of the CN from the so-called normal behavior profile (NBP). Although a lot of scientific and practical works are devoted to the development of NBP, however, successful attempts of network attacks on a number of domestic and foreign institutions indicate the need for their significant improvement, which explains the relevance of research in this direction.

2 Literature Review and Problem Postulation

As evidenced by the literature review, control cards Shuhart, EWMA, CUSUM, simulation and neural network models mainly used for the construction of modern PNP, as well as models based on classical methods of statistical analysis [5, 6, 18, 19]. It should be noted that in most cases, these profiles describe the dynamics of the functional parameters of the CN. The method and experimental system of detection of attacks on resources of CN is developed in work [9]. The method uses the patterns of normal behavior based on the simulation model of the functioning of network objects of CN. It should be mentioned that the construction of this simulation model for real CS is associated with great difficulties. The aim of the work [1] was to develop a mathematical model of CS network behavior and a method of anomaly detection based on the statistical study of changes in the characteristics of this system. This paper proposes to divide the entire space of the performance on the indicators characterizing the use of local system resources and indicators of interaction with the outside world. The introduced mathematical model based on the assumption the use of the sample homogeneity criterion to identify anomalies. In addition, the literature review allows us to determine that the methods of statistical data processing can be additionally used to clarify the obtained NBP. Thus, the article [4] describes the process of developing templates for fixing abnormal behavior of CN on the basis of BDS-tests. It is argued that such tests are effective methods for identifying dependencies in time series in the framework of their nonlinear analysis and have been widely used for analysis of financial markets. In the article the mathematical apparatus of testing is developed, and the results of numerical experiments are shown, confirming the possibility of using such templates to define the malicious software by heuristic analyzers. Article [13] is devoted to the development of adaptive patterns of fixation of abnormal behavior of CN. The fact of possibility of use for these purposes of control cards of Shuhart is confirmed. For the development and experimental studies, a software model has been developed that allows to obtain a database of CN state templates and to fix abnormal behavior of network objects. The possibility of using the developed adaptive patterns in heuristic analyzers of intrusion detection systems is stated. However, as studies have shown, these means of forming patterns are not devoid of a number of drawbacks. SO EWMA control cards are insensitive to short manifestations of anomalies. At the same time, CUSUM maps detect small but constant changes with a higher probability but have low accuracy (high probability of false positives) under dynamic changes in the computer system display [2-5]. The analyzed patterns, which are based on classical methods of statistical analysis, are characterized by a high level of false positives when used in local networks, where the behavior of objects does not have a smooth, averaged character [8, 10]. Based on authors view, it is possible to eliminate these shortcomings by using the theory of Markov processes, which is successfully used to build statistical models of complex processes [7]. It should be noted that the expediency of using Markov chains in the field of recognition of cyberattacks has already been proved in the works [9, 19, 20]. Thus, the work [9] is devoted to the use of hidden Markov chains for modeling the mental state of the CN operator in the implementation of cyberattacks. The paper [14] describes in detail the methodology of CN protection based on the game theory using the Markov model. The possibility of choosing the most effective protection strategy is shown in the pre-cases of a certain time window. A similar problem is considered in work [15], where the Markov model is developed to determine the strategy of protection of a moving target. In work [17] the algorithm of step-by-step detection of attacks on a computer network on the basis of the hidden Markov model is offered. In work [20] the possibility of formation of templates of normal behavior of network objects of CN on the basis of a homogeneous Markov chain with consecutive transitions is proved. It is determined that for the formation of patterns of normal behavior it is advisable to use a Markov chain with the number of States equal to 20. The graph of process transitions is developed, the corresponding mathematical support allowing to calculate the basic parameters of the Markov link underlying the specified template is formed. The results are experimental, confirming the effectiveness of the developed Markov model for conditionally stationary dynamics of controlled parameters of protection. Moreover, the expediency of further studies in the field of justification of the nomenclature of controlled parameters and adaptation of the Markov model of the typical non-stationary nature of the functional parameters of the CN is shown. It should be noted that the possibility of creating a Markov model of unsteady NBP is confirmed in theoretical works devoted to the Markov approximation of multiperiodic unsteady processes [7, 19].

Thus, as a result of the analysis of scientific and practical works [2-4, 7-17, 19, 20], the prospects of using Markov chains for the formation of NBP are determined. In addition, it is shown that the known mathematical models of NBP do not fully consider the typical non-stationary nature of the dynamics of the functional parameters of network CN.

Purposes and Objectives of the Study. The main purpose of this study is to develop a Markov model of the non-stationary profile of the normal behavior of network objects of computer systems, which is adapted to the typical non-stationary nature of the dynamics of the auxiliary parameters of such objects.

3 Development of the Markov model

In the construction of the Markov model, the position of the theory of dynamic data series is used, according to which the multiperiodic NBP can be represented as a sum of single-period profiles. The graph of single-period NBP is shown in Fig. 1, and the graph of multiperiodic PNP is shown in Fig. 2.

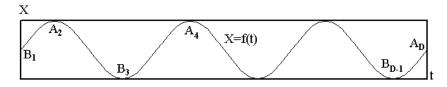


Fig. 1. Single-period NBP graph.

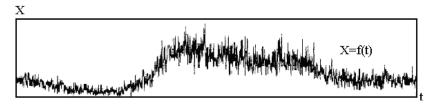


Fig. 2. Graph of multiperiodic NBP.

The actual NBP corresponds to the function X=f(t). In Fig. 1, the letters A and B denote the transition (extreme) points of the function X=f (t)., the maxima of the function are denoted as A₂, A₄, ... A_D, and the minima as B₁, B₂, ... B_{D-1}. Indices 1,2, ... D correspond to the numbers of transition points. At intervals of type $B_d A_{d+1}$ and

 $A_{d+1}B_{d+2}$ single-period NBP has a stationary character. On intervals of type $B_d A_{d+1}$ the function X=f (t) increases, and on intervals of type $A_{d+1}B_{d+2}$ the function X=f (t) decreases.

According to theoretical studies in the field of the theory of Markov processes, the basis of NBP on a stationary interval of the dynamics of operational parameters of CN is a homogeneous Markov chain described by a system of Kolmogorov-Chapman equations and a normalization condition that can be written using expressions:

$$\begin{cases} P_{1}(t) = P_{1}(t-1) - P_{1}(t-1)p_{1,i} \dots - P_{1}(t-1)p_{1,N} + P_{i}(t-1)p_{i,1} \dots + P_{N}(t-1)p_{N,1} \\ \dots \\ P_{i}(t) = P_{i}(t-1) - P_{i}(t-1)p_{i,1} \dots - P_{i}(t-1)p_{i,N} + P_{1}(t-1)p_{1,i} \dots + P_{N}(t-1)p_{N,i} \\ \dots \\ P_{N}(t) = P_{N}(t-1) - P_{N}(t-1)p_{N,1} \dots - P_{N}(t-1)p_{N,i} + P_{1}(t-1)p_{1,N} \dots + P_{i}(t-1)p_{i,N} \\ \sum_{i=1}^{N} P_{i}(t) = 1 \end{cases}$$

$$(2)$$

where $P_i(t)$ is the probability of finding the functional parameter in the i-th state at time t $\in [0, t_{max}]$, $p_{i,j}$ - is the probability of transition from state i to state j in one step of the process, N is the number of States of the Markov chain.

If we accept the postulate that at the initial moment of time the simulated parameter is in the first state of the Markov chain, then the initial conditions of modeling can be written as follows:

$$P_1(0) = 1$$
 (3)

The disadvantage of the described Markov chain is the complexity of the calculation of transition probabilities, which is determined primarily by the fully connected character of possible transitions of the process by states.

Therefore, based on the results of [7, 8, 19], the assumption is made about the possibility of using a Markov chain in which only successive transitions are possible between states, the number of which is equal to 10. At the same time, the accuracy of the model remains sufficient for the task of forming the NBP. Due to the accepted simplification, expression (1) is modified as follows:

$$\begin{cases}
P_{1}(t) = P_{1}(t-1) - P_{1}(t-1)p_{1,2} + P_{2}(t-1)p_{2,1} \\
\dots \\
P_{i}(t) = P_{i}(t-1) - P_{i}(t-1)p_{i,i+1} + P_{i+1}(t-1)p_{i+1,i} , \\
\dots \\
P_{N}(t) = P_{N}(t-1) - P_{N}(t-1)p_{N,N-1} + P_{N-1}(t-1)p_{N-1,N}
\end{cases}$$
(4)

where N = 10 is the number of states of the Markov chain.

In case of approximation of dynamics of operational parameters of computer networks it is recommended that the states with numbers from 1 to (N-1) correspond to such values of parameters at which the software and hardware of a network will be in a working state, and the N-th state corresponds to a non-working state [7]. In a one-sided region of operability, the first (N-1) states will be defined by the lower and upper bounds, and the last N state by the lower bound only. Using the procedure of uniform quantization of the region of operability, the state boundaries are defined as follows:

$$L_i^b = \Delta \times i \,, \tag{5}$$

$$L_i^e = \Delta \times (i-1), \tag{6}$$

$$\Delta = \frac{L}{N-1},\tag{7}$$

$$L = L^b - L^e, (8)$$

where, L_i^b , L_i^e is the upper and lower bound of the i-th state, Δ -is the width of the state, L-- width of the region of operability, L^b , L^e - upper and lower limit of the region of operability.

An illustration of the described quantization procedure is Fig. 3, which shows graphs of the dynamics of some operational parameter X for the same type of controlled objects. In Fig. 3, the following notations are accepted: t_k – moments of registration; 1...N – numbers of states of a Markov chain; n_i – borders of states; O_1 , O_2 , O_3 , O_4 , O_5 -realizations of X. The calculation of the probabilities of transition of the controlled parameter from state i to state j in one step of the process is implemented as follows:

$$p(t_k)_{i=1,i} = R(t_k)_i / R(t_k)_r,$$
(9)

$$p_{i-1,i} = \sum_{i=1}^{I} p(t_k)_{i-1,i} / I , \qquad (10)$$

where $p(t_k)_{i-1,i}$ – is the probability of transition between (i-1)-th and i-th state in one step of calculation of Markov chains $R(t_k)_i$ – the number of controlled objects that

are moved from (i-1)-th to i-th state for the time interval t_k , which corresponds to one step of the Markov chain. R $(t_k)_r$ — the total number of operable objects at the time t_k , $p_{i-1,i}$ -the probability of transition between (i-1)-th and i-th state, I-the number of registrations of the parameter.

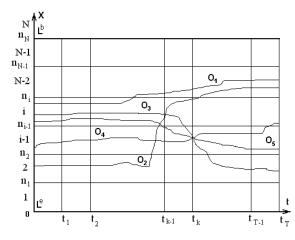


Fig. 3. The graphs of dynamics of values of operational parameter for controlled objects

In the simplest case, to simulate one period of a single-period process of changing the operational parameter X, the graph of which is shown in Fig. 3, the model will consist of two homogeneous Markov chains given by expressions (2, 3, 4). The first circuit is designed to simulate the AB section, the second-for the BC section. In this case, the probabilities of the first Markov chain at time B are the initial distribution of the form (1) for the second Markov chain.

For a two-period process, the graph of which is shown in Fig. 4, the Mars model will be more complex. In such a model, it is necessary to consider that the BF section is non-stationary only in the CE section. In this case, the half-period CD and DE are nested in the half-period BF.

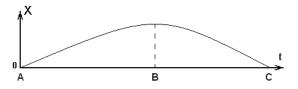


Fig. 3. The half-life graph of a single-period process.

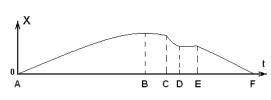


Fig. 4. The graph of two-period process.

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To construct a multiperiodic Markov model, the authors developed an approach based on the Markov model described in [19, 20] of a one-periodic nonstationary NBP process X = f(t), which increases sequentially at stationary intervals of type $B_d A_{d+1}$ and decreases at stationary intervals of type $A_{d+1} B_{d+2}$, where d is the number of the transition point (see Fig. 1). The Markov model of one-period NBP M_{BAB} consists of two homogeneous Markov chains M_{BA} and M_{AB} , designed to simulate the dynamics of functional parameters at stationary intervals of type $B_d A_{d+1}$ and $A_{d+1}B_{d+2}$. By analogy with [2, 3] in the model of multiperiodic NBP, it is also assumed that each of the stationary sites is modeled by its own homogeneous Markov chain. For example, the NBP of a two-period process, the graph of which is shown in fragments in Fig. 4, a sequential simulation of each of the stationary intervals AB, BC, CD, DE, and EF is provided by a proper homogeneous Markov chain. As for the case of a one-period NBP, the finite probability distribution of the stay of the previous Markov chain is the initial conditions for the subsequent chain. Thus, the Markov model of multiperiodic NBP M_{BAB}^{Σ} , the structure of which is shown in Fig. 5, consists of modules $M_{BAB}^{(1)}, M_{BAB}^{(2)}, ..., M_{BAB}^{(K)}$ designed to simulate K significant periods of the dynamics of the functional parameter under study. In this case, an arbitrary Kth module $M_{BAB}^{(k)}$ is a Markov model of a one-period NBP developed in [19, 20], designed for modeling the k-th periodic component. In turn $M_{BAB}^{(k)}$, it consists of two LM- $M_{RA}^{(k)}$ and $M_{AR}^{(k)}$, designed to model the k-th periodic component of the NBP... The output of the k-th module $M^{(k)}_{BAB}$ at the τ -th step of the calculation is $\|P^{(k)}(\tau)\|$ the probability distribution vector for the k-th periodic component of the NBP. The number of states and state boundaries of the Markov chain of each of the modules should be calculated individually using the expressions (5-8). Also, the values of transition probabilities should be calculated separately for each Markov chain. For this purpose, it is possible to use expressions (9, 10), having modified them a little for processing of statistics on the corresponding interval of functioning (half-life):

$$p_{z}(t_{k})_{i-1,i} = R_{z}(t_{k})_{i} / (R_{z}(t_{k})_{r} \times t_{k}),$$
(11)

$$p_{z,i-1,i} = \sum_{i=1}^{I_z} p_z(t_k)_{i-1,i} / I_z , \qquad (12)$$

where z is the half-period in question.

Other symbols (11, 12) correspond to expressions (9, 10).

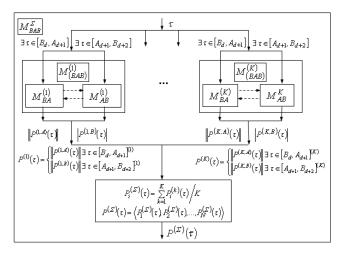


Fig. 5. Markov model of multiperiodic NBP

The output of the model M_{BAB}^{Σ} at the τ t-th step of calculation is the probability distribution vector of the form:

$$\left\|P^{(\Sigma)}(\tau)\right\| = \left\langle P_1^{(\Sigma)}(\tau), P_2^{(\Sigma)}(\tau), \dots, P_N^{(\Sigma)}(\tau)\right\rangle,$$
(13)

where $P_i^{(\Sigma)}(\tau)$ is the integral probability of finding the parameter in the i-th state of the Markov chain at the τ -th step of the calculation.

 $P_i^{(\Sigma)}(\tau)$ it is calculated as follows:

$$P_{i}^{(\Sigma)}(\tau) = K^{-1} \sum_{k=1}^{K} P_{i}^{(k)}(\tau), \qquad (14)$$

where $P_i^{(k)}(\tau)$ is the probability of the functional parameter staying in the i-th state of the K-th Markov chain at the τ -th calculation step.

Using the proposed Markov model, the software package MarkPr was developed, which allows to simulate Markov processes of various types. So, in Fig.6 shows the simulation results of a two-period Markov process with 10 states. The graph shown in Fig. 6. corresponds to the mathematical expectation of the controlled parameter.

0 0.051470 0.060279	0.019203	0.075375	0.020163	0.075826
0.051470				
0.060279	0.027086	0.080582	0.000000	
			0.028232	0.080944
0.068633	0.036224	0.085159	0.037477	0.085439
0.077477	0.047307	0.089483	0.048554	0.089683
0.090408	0.062160	0.096181	0.063254	0.096258
0.103736	0.081699	0.102908	0.082449	0.102861
0.120289	0.110851	A 0.110150	0.110813	0.109962
0.129133	0.140326	0.114485	0.139834	0.114216
0.139297	0.189004	B 0.119453	0.187542	0.119089
0.159276	0.286139	0.126225	0.281683	0.125721
	0.077477 0.090408 0.103736 0.120289 0.129133 0.139297 0.159276	0.077477 0.047307 0.090408 0.052160 0.103736 0.081699 0.120289 0.110051 0.129133 0.140326 0.139297 0.189004 0.159276 0.268139	0.077477 0.047307 0.089483 0.090408 0.052160 0.096181 0.103736 0.081699 0.102308 0.120289 0.110851 0.110150 0.129133 0.140326 0.114485 0.139297 0.189004 0.119453 0.139297 0.189004 0.119453 0.159276 0.206139 0.128225	0.077477 0.047307 0.089483 0.048554 0.090408 0.062160 0.096181 0.063254 0.103736 0.081699 0.102008 0.082449 0.120209 0.110851 0.110150 0.110813 0.129133 0.140326 0.114485 0.139344 0.139297 0.189004 0.0119453 0.167542

Fig. 6. The graphs of the dynamics of the mathematical expectation of the number of requests

On the basis of the described Markov model, given by the expressions (2-14), a two-period NBP of a Web server is developed. Statistical data were used to construct the model [19, 20]. The simulation was carried out using the mentioned MarkPr program. As a security parameter, the number of web server accesses per 1 minute is used. The simulation results are partially presented in Fig.7, on which 1 denotes a graph based on statistical data, 2 If based on a one-period model [20], and 3 based on the proposed two-period model.

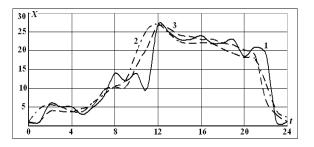


Fig. 7. The graphs of the dynamics of the mathematical expectation of the number of requests

It is important to mention that for a one-period model, the average modeling error is 0.09 [19, 20], for a two - period (author's) model-0.04, and for common polynomial models [1, 4] - 0,14 \div 0,18. Therefore, the application of the proposed two-period model allowed to reduce the error of modeling with respect to the one-period model by about 2 times, and with respect to common polynomial models-by about 4 times. Consequently, the results of the experiments confirm the effectiveness of the pro-

posed two-period Markov model of NBP. At the same time, the issue of determining the number of significant periods that need to be taken into account in the NBP remains unresolved. Based on the results [12], it can be assumed that this deficiency can be corrected by using the theory of spectral analysis of data.

4 Conclusion

As a result of the conducted researches, it is justified expediency forming of multiperiodic profiles of normal behavior of network objects of computer systems on the basis of a Markov chain with consecutive transitions *is proved*. The developed Markov model of the non-stationary profile of normal behavior allows to consider the typical bagatoperiodic nature of the dynamics of functional parameters , which reflect the state of security of network objects of computer systems. As a result of experimental studies, it was found that the application of the developed multiperiodic model allowed to increase the accuracy of forecasting the dynamics of functional parameters up to 2 times relative to the best known models of similar purpose. It is shown that the prospects for further research are associated with the development of methods for applying solutions of the theory of spectral analysis of data to determine the significant periods of the process of change of functional parameters.

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